

# Modified Generalized Chain Ratio in Regression Type Estimator for Population Mean using Two Phase Sampling in the Presence of Non-Response

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**Abstract:** Modified generalized chain ratio in regression type estimator for population mean using two phase sampling in the presence of non-response has been proposed and their properties have been studied. Relative efficiency of the proposed estimator is obtained in the case of fixed first phase sample, second phase sample and sub sample fraction. A comparison of the proposed estimator has been carried out with the relevant estimators. The performances of the proposed estimator in comparison to the relevant estimators have been made with the help of an empirical study.

**Keywords:** Non-response; Mean square error; two phase sampling; Auxiliary characters.

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## 1 Introduction

The information on the auxiliary character provides a very important contribution in the field of sample surveys. In the case when the population mean of the auxiliary character  $x$  is known, Cochran [2], Rao [3,4,5] and Khare and Srivastava [6,7] have proposed ratio, product and regression type estimators in the presence of non-response. In the case when the population mean of an auxiliary character is not known, the two phase sampling ratio, product and regression type estimators have suggested by Khare and Srivastava [8,9], Khare et al. [10], Khare and Kumar [11] and Khare and Srivastava [12].

In such case, Chand [13] proposed chain ratio type estimator for the population mean of the study character. Kiregyra [14,15] proposed chain ratio to regression, ratio in regression and regression in regression estimators for the population mean of the study character. Further, some class of estimators for population mean of the study character have been proposed by Srivastava et al. [16], Sahoo et al. [17], Singh et al. [18] and Dash & Mishra [19].

In the case when the population mean of the main auxiliary character is not known but the population mean of an additional auxiliary character is known, which is cheaper than the main auxiliary character but less correlated to the study character than the main auxiliary character, Khare and Kumar [20] and Khare et al. [21] have proposed chain regression type estimators and generalized chain estimators for the population mean in the presence of non-response.

In the present paper, we have proposed modified generalized chain ratio in regression type estimator for the population mean using two phase sampling in the presence of non-response. We have obtained the expressions for mean square error of the proposed estimator for the fixed first phase sample, second phase sample and the optimum value of constants. A comparative study of the variance of the proposed estimator is made with the relevant estimators. An empirical study has been given to show the performance of the proposed estimator for fixed sample sizes.

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## 2 The Estimators

Let  $\bar{Y}, \bar{X}$  and  $\bar{Z}$  denote the population mean of study character  $y$ , auxiliary character  $x$  and additional auxiliary character  $z$  having  $j^{\text{th}}$  value  $Y_j, X_j$  and  $Z_j; j = 1, 2, \dots, N$ . Population of size  $N$  is supposed to be divided in  $N_1$  responding units and  $N_2$  non respondents units. According to Hansen and Hurwitz [1], a sample of size  $n$  is drawn from the population of size  $N$  by using simple random sampling without replacement (SRSWOR) method of sampling and it has been observed that  $n_1$  units respond and  $n_2$  units do not respond. Again by making extra effort, a sub-sample of size  $r (= n_2 k^{-1})$  is drawn from  $n_2$  non-responding units by using SRSWOR method of sampling and collect information on  $r$  units for study character  $y$ . Hence the estimator for  $\bar{Y}$  based on  $n_1 + r$  observations on  $y$  is given by Hansen and Hurwitz[1] as follow:

$$\bar{y}^* = \frac{n_1}{n} \bar{y}_1 + \frac{n_2}{n} \bar{y}'_2 \quad (1)$$

where  $\bar{y}_1$  and  $\bar{y}'_2$  are the sample means of study character  $y$  based on  $n_1$  and  $r$  units respectively. Using the values of  $x$  corresponding to the incomplete information on  $y$ , the estimator for population mean  $\bar{X}$  of auxiliary character  $x$  is defined as follow:

$$\bar{x}^* = \frac{n_1}{n} \bar{x}_1 + \frac{n_2}{n} \bar{x}'_2 \quad (2)$$

Using the values of  $z$  corresponding to the incomplete information on  $y$ , the estimator for population mean  $\bar{Z}$  of additional auxiliary character  $z$  is defined as follow:

$$\bar{z}^* = \frac{n_1}{n} \bar{z}_1 + \frac{n_2}{n} \bar{z}'_2 \quad (3)$$

Variance of the estimators  $\bar{y}^*, \bar{x}^*$  and  $\bar{z}^*$  are given by

$$V(\bar{y}^*) = \frac{f}{n} S_y^2 + \frac{W_2(k-1)}{n} S_{y(2)}^2 \quad (4)$$

$$V(\bar{x}^*) = \frac{f}{n} S_x^2 + \frac{W_2(k-1)}{n} S_{x(2)}^2 \quad (5)$$

$$V(\bar{z}^*) = \frac{f}{n} S_z^2 + \frac{W_2(k-1)}{n} S_{z(2)}^2 \quad (6)$$

where  $f = 1 - \frac{n}{N}, W_2 = \frac{N_2}{n}, (S_y^2, S_{y(2)}^2), (S_x^2, S_{x(2)}^2)$  and  $(S_z^2, S_{z(2)}^2)$  are population mean squares of  $y, x$  and  $z$  for entire population and non-responding part of population.

In the case when population mean  $N$  is not known, a first phase sample of size  $n' (< N)$  is taken from the population of size  $N$  by using SRSWOR scheme of sampling and the population mean  $\bar{X}$  is estimated by first phase sample mean  $\bar{x}'$  based on  $n'$  units. In this case, the conventional / alternative generalized two phase sampling estimators  $(T_1, T_2)$  for population mean  $\bar{Y}$  in the presence of non-response have been proposed by Khare and Srivastava [12], which are given as follows:

$$T_1 = \bar{y}^* \left( \frac{\bar{x}^*}{\bar{x}'} \right)^\alpha \quad (7)$$

$$T_2 = \bar{y}^* \left( \frac{\bar{x}}{\bar{x}'} \right)^\beta \quad (8)$$

where  $\bar{x}' = \frac{1}{n'} \sum_{j=1}^{n'} x_j, \bar{x} = \frac{1}{n} \sum_{j=1}^n x_j$  and  $(\alpha, \beta)$  are constants.

In the situation when population mean  $\bar{X}$  of the auxiliary character  $x$  is unknown, but  $\bar{Z}$ , the population mean of additional auxiliary character  $z$  (closely related to  $x$ ) is known, which may be cheaper and less correlated to the study character  $y$  in comparison to main auxiliary character  $x$ , a first phase sample of size  $n' < N$  is taken from the population of size  $N$  by using SRSWOR scheme of sampling and the information of auxiliary character  $x$  and additional auxiliary character  $z$  based on  $n'$  units is collected. Using the sample means  $\bar{x}'$  and  $\bar{z}'$  based on  $n'$  units and the known additional population mean  $\bar{Z}$ , the population mean  $\bar{X}$  is estimated by  $\hat{X}_r = \bar{x}' \bar{Z} / \bar{z}'$ . The estimator  $\hat{X}_r$  is more precise than sample mean  $\bar{x}'$  if  $\rho_{xz} > C_z / 2C_x$ , where  $C_z, C_x$  and  $\rho_{xz}$  are the coefficients of variation of  $z$  and  $x$  and the correlation coefficient between  $x$  and  $z$ .

In this situation, the conventional  $(T_3)$  and alternative  $(T_4)$  generalized chain estimators for  $\bar{Y}$  using information on auxiliary character  $x$  and the additional auxiliary character  $z$  in the presence of non-response have been proposed by Khare et al. [12], which are given as follows:

$$T_3 = \bar{y}^* \left( \frac{\bar{x}^*}{\bar{x}} \right)^{\alpha_1} \left( \frac{\bar{z}'}{\bar{z}} \right)^{\alpha_2} \tag{9}$$

$$T_4 = \bar{y}^* \left( \frac{\bar{x}}{\bar{x}^*} \right)^{\beta_1} \left( \frac{\bar{z}'}{\bar{z}} \right)^{\beta_2} \tag{10}$$

where  $\alpha_1, \alpha_2, \beta_1$  and  $\beta_2$  are constants.

In the case, when both population means  $\bar{X}$  and  $\bar{Z}$  are unknown, complete information on the auxiliary variables is collected for all the units in the first-phase sample. The collected information becomes available for sub sampling phases. We proposed modified generalized chain ratio in regression type estimator for population mean using two phase sampling in the presence of non-response, which is given as follow:

$$T_5 = \bar{y}^* + b_{yx} \left\{ \bar{x}' \left( \frac{\bar{z}'}{\bar{z}^*} \right)^{p_1} \left( \frac{\bar{z}}{\bar{z}'} \right)^{p_2} - \bar{x}^* \right\} \tag{11}$$

where  $p_1$  and  $p_2$  are constants and  $b_{yx}$  is regression coefficient.

### 3 Mean Square Error of the Proposed Estimator

In order to derive the expressions for mean square error of the estimators-

Let  $\bar{y}^* = \bar{Y}(1 + \epsilon_0)$ ,  $\bar{x}^* = \bar{X}(1 + \epsilon_1)$ ,  $\bar{x}' = \bar{X}(1 + \epsilon_2)$ ,  $\bar{z}^* = \bar{Z}(1 + \epsilon_3)$ ,  $\bar{z}' = \bar{Z}(1 + \epsilon_4)$  and  $\bar{z} = \bar{Z}(1 + \epsilon_5)$ , such that  $E(\epsilon_l) = 0$  and  $\epsilon_l < 1$ ,  $l = 0, 1, 2, 3, 4, 5$ . Now, using SRSWOR scheme of sampling, we have

$$\begin{aligned} E(\epsilon_0^2) &= \frac{f}{n} C_y^2 + \frac{W_2(k-1)}{n} C_{y(2)}^2, & E(\epsilon_1^2) &= \frac{f}{n} C_x^2 + \frac{W_2(k-1)}{n} C_{x(2)}^2, & E(\epsilon_2^2) &= \frac{f'}{n'} C_x^2, \\ E(\epsilon_3^2) &= \frac{f}{n} C_z^2 + \frac{W_2(k-1)}{n} C_{z(2)}^2, & E(\epsilon_4^2) &= \frac{f'}{n'} C_z^2, & E(\epsilon_5^2) &= \frac{f}{n} C_z^2, \\ E(\epsilon_0 \epsilon_1) &= \frac{f}{n} C_{yx} + \frac{W_2(k-1)}{n} C_{yx(2)}, & E(\epsilon_0 \epsilon_2) &= \frac{f'}{n'} C_{yx}, & E(\epsilon_0 \epsilon_3) &= \frac{f}{n} C_{yz} + \frac{W_2(k-1)}{n} C_{yz(2)}, \\ E(\epsilon_0 \epsilon_4) &= \frac{f'}{n'} C_{yz}, & E(\epsilon_0 \epsilon_5) &= \frac{f}{n} C_{yz}, & E(\epsilon_1 \epsilon_2) &= \frac{f'}{n'} C_x^2, & E(\epsilon_1 \epsilon_3) &= \frac{f}{n} C_{xz} + \frac{W_2(k-1)}{n} C_{xz(2)}, \\ E(\epsilon_1 \epsilon_4) &= \frac{f'}{n'} C_{xz}, & E(\epsilon_1 \epsilon_5) &= \frac{f}{n} C_{xz}, & E(\epsilon_2 \epsilon_3) &= \frac{f'}{n'} C_{xz}, & E(\epsilon_2 \epsilon_4) &= \frac{f'}{n'} C_{xz}, \\ E(\epsilon_2 \epsilon_5) &= \frac{f}{n} C_{xz}, & E(\epsilon_3 \epsilon_4) &= \frac{f'}{n'} C_z^2, & E(\epsilon_3 \epsilon_5) &= \frac{f}{n} C_z^2, & E(\epsilon_4 \epsilon_5) &= \frac{f'}{n'} C_z^2 \end{aligned}$$

The contribution of the terms involving the powers in  $\epsilon_0, \epsilon_1, \epsilon_2, \epsilon_3, \epsilon_4$  and  $\epsilon_5$  of order higher than two in mean square error are assumed to be negligible. Hence, the expression for mean square error of the estimator  $T_5$  upto the terms of order  $(1/n)$  is given by

$$\begin{aligned} MSE(T_5) &= V(\bar{y}^*) + \lambda b \bar{X} [b \bar{X} C_z^2 + p_1^2 b \bar{X} C_z^2 - 2 \bar{Y} C_{yx} - 2 p_1 \bar{Y} C_{yz} + 2 p_1 b \bar{X} C_{xz} + 2 p_1 p_2 \bar{Y} C_{yz} - 2 p_2 b \bar{X} C_{xz} \\ &\quad + p_2^2 b \bar{X} C_z^2 - 2 p_1 p_2 b \bar{X} C_z^2] + \theta b \bar{X} [b \bar{X} C_{z(2)}^2 + p_1^2 b \bar{X} C_{z(2)}^2 - 2 \bar{Y} C_{yx(2)} - 2 p_1 \bar{Y} C_{yz(2)} + 2 p_1 b \bar{X} C_{xz(2)} \\ &\quad + 2 p_1 p_2 \bar{Y} C_{yz(2)} - 2 p_2 b \bar{X} C_{xz(2)} + p_2^2 b \bar{X} C_{z(2)}^2 - 2 p_1 p_2 b \bar{X} C_{z(2)}^2] \end{aligned} \tag{12}$$

where  $p_1$  and  $p_2$  are optimum values and  $b$  is regression coefficient.

$$p_{1(opt)} = \frac{\lambda (\bar{Y}C_{yz} - b\bar{X}C_{xz} - p_2b\bar{X}C_z^2 + p_2\bar{Y}C_{yz}) + \theta (\bar{Y}C_{yz(2)} - b\bar{X}C_{xz(2)})}{b\bar{X} (\lambda C_z^2 + \theta C_{z(2)}^2)} \quad (13)$$

$$p_{2(opt)} = \frac{b\bar{X}C_{xz} - \bar{Y}C_{yz} + Mb\bar{X}C_z^2}{b\bar{X}C_z^2 - Qb\bar{X}C_z^2} \quad (14)$$

$$\text{where, } \lambda = \left( \frac{1}{n} - \frac{1}{n'} \right), \quad \theta = \frac{W_2(k-1)}{n}, \quad b = \frac{\bar{Y}\rho_{yx(2)}C_{y(2)}}{\bar{X}C_{x(2)}} \quad (15)$$

$$M = \frac{\lambda (\bar{Y}C_{yz} - b\bar{X}C_{xz}) + \theta (\bar{Y}C_{yz(2)} - b\bar{X}C_{xz(2)})}{b\bar{X} (\lambda C_z^2 + \theta C_{z(2)}^2)}, \quad Q = \frac{\lambda C_z^2}{(\lambda C_z^2 + \theta C_{z(2)}^2)} \quad (16)$$

Mean square error of the estimators  $T_1, T_2, T_3$  and  $T_4$  upto the terms of order  $(1/n)$  is given by

$$MSE(T_1) = V(\bar{y}^*) - \bar{Y}^2 \left[ \lambda \rho_{yx}^2 C_y^2 + \theta \rho_{yx(2)}^2 C_{y(2)}^2 \right] \quad (17)$$

$$MSE(T_2) = V(\bar{y}^*) - \bar{Y}^2 \left[ \lambda \rho_{yx}^2 C_y^2 \right] \quad (18)$$

$$MSE(T_3) = V(\bar{y}^*) - \bar{Y}^2 \left[ \lambda \rho_{yx}^2 C_y^2 + \theta \rho_{yx(2)}^2 C_{y(2)}^2 + \eta \rho_{yz}^2 C_y^2 \right] \quad (19)$$

$$MSE(T_4) = V(\bar{y}^*) - \bar{Y}^2 \left[ \lambda \rho_{yx}^2 C_y^2 + \eta \rho_{yz}^2 C_y^2 \right] \quad (20)$$

where,

$$C_{yx} = \rho_{yx} C_y C_x, \quad C_{yx(2)} = \rho_{yx(2)} C_{y(2)} C_{x(2)}, \quad C_{yz} = \rho_{yz} C_y C_z, \\ C_y = S_y/\bar{Y}, \quad C_x = S_x/\bar{X}, \quad C_z = S_z/\bar{Z}, \quad C_{y(2)} = S_{y(2)}/\bar{Y}, \quad C_{x(2)} = S_{x(2)}/\bar{X},$$

$$C_{z(2)} = S_{z(2)}/\bar{Z}, \quad S_y^2 = \frac{1}{(N-1)} \sum_{j=1}^N (Y_j - \bar{Y})^2, \quad S_x^2 = \frac{1}{(N-1)} \sum_{j=1}^N (X_j - \bar{X})^2,$$

$$S_z^2 = \frac{1}{(N-1)} \sum_{j=1}^N (Z_j - \bar{Z})^2, \quad S_{y(2)}^2 = \frac{1}{(N_2-1)} \sum_{j=1}^{N_2} (Y_j - \bar{Y}_2)^2,$$

$$S_{x(2)}^2 = \frac{1}{(N_2-1)} \sum_{j=1}^N (X_j - \bar{X}_2)^2, \quad S_{z(2)}^2 = \frac{1}{(N_2-1)} \sum_{j=1}^{N_2} (Z_j - \bar{Z}_2)^2, \quad \eta = \frac{f'}{n'}$$

$(\rho_{yx}, \rho_{yx(2)})$  are the correlation coefficients between  $(y, x)$  for whole population and non-responding part of the population,  $(\rho_{yz}, \rho_{yz(2)})$  is the correlation coefficient between  $(y, z)$  for whole population and  $(\bar{Y}_2, \bar{X}_2)$  are the means of the study character  $y$  and auxiliary character  $x$  and auxiliary character  $z$  for the non-responding part of the population.

### 4 Comparison of the Proposed Estimator $T_5$ with Respect to the Relevant Estimators

$$\begin{aligned}
 &MSE(T_5) < MSE(\bar{y}^*) \\
 \text{if } &\frac{-A_2 - \sqrt{A_2^2 - 4A_1A_3}}{2A_1} < p_1 < \frac{-A_2 + \sqrt{A_2^2 - 4A_1A_3}}{2A_1} \\
 \text{and } &\frac{-B_2 - \sqrt{B_2^2 - 4B_1B_3}}{2B_1} < p_2 < \frac{-B_2 + \sqrt{B_2^2 - 4B_1B_3}}{2B_1} \tag{21}
 \end{aligned}$$

$$\begin{aligned}
 &MSE(T_5) < MSE(T_1) \\
 \text{if } &\frac{-A_2 - \sqrt{A_2^2 - 4A_1A_4}}{2A_1} < p_1 < \frac{-A_2 + \sqrt{A_2^2 - 4A_1A_4}}{2A_1} \\
 \text{and } &\frac{-B_2 - \sqrt{B_2^2 - 4B_1B_4}}{2B_1} < p_2 < \frac{-B_2 + \sqrt{B_2^2 - 4B_1B_4}}{2B_1} \tag{22}
 \end{aligned}$$

$$\begin{aligned}
 &MSE(T_5) < MSE(T_2) \\
 \text{if } &\frac{-A_2 - \sqrt{A_2^2 - 4A_1A_5}}{2A_1} < p_1 < \frac{-A_2 + \sqrt{A_2^2 - 4A_1A_5}}{2A_1} \\
 \text{and } &\frac{-B_2 - \sqrt{B_2^2 - 4B_1B_5}}{2B_1} < p_2 < \frac{-B_2 + \sqrt{B_2^2 - 4B_1B_5}}{2B_1} \tag{23}
 \end{aligned}$$

$$\begin{aligned}
 &MSE(T_5) < MSE(T_3) \\
 \text{if } &\frac{-A_2 - \sqrt{A_2^2 - 4A_1A_6}}{2A_1} < p_1 < \frac{-A_2 + \sqrt{A_2^2 - 4A_1A_6}}{2A_1} \\
 \text{and } &\frac{-B_2 - \sqrt{B_2^2 - 4B_1B_6}}{2B_1} < p_2 < \frac{-B_2 + \sqrt{B_2^2 - 4B_1B_6}}{2B_1} \tag{24}
 \end{aligned}$$

$$\begin{aligned}
 &MSE(T_5) < MSE(T_4) \\
 \text{if } &\frac{-A_2 - \sqrt{A_2^2 - 4A_1A_7}}{2A_1} < p_1 < \frac{-A_2 + \sqrt{A_2^2 - 4A_1A_7}}{2A_1} \\
 \text{and } &\frac{-B_2 - \sqrt{B_2^2 - 4B_1B_7}}{2B_1} < p_2 < \frac{-B_2 + \sqrt{B_2^2 - 4B_1B_7}}{2B_1} \tag{25}
 \end{aligned}$$

where,

$$\begin{aligned}
 A_1 &= b\bar{X}(\lambda C_z^2 + \theta C_{z(2)}^2), & A_2 &= \lambda(-2\bar{Y}C_{yz} + 2b\bar{X}C_{xz} - 2Pb\bar{X}C_z^2) + \theta(-2\bar{Y}C_{yz(2)} + 2b\bar{X}C_{xz(2)}), \\
 A_3 &= \lambda(b\bar{X}C_x^2 - 2\bar{Y}C_{yx} + 2Pb\bar{Y}C_{yz} + 2Pb\bar{X}C_{xz} - b\bar{X}PC_z^2) + \theta(b\bar{X}C_{x(2)}^2 - 2\bar{Y}C_{yx(2)}), \\
 B_1 &= \lambda b\bar{X}C_z^2(1-Q)^2 + \theta Q^2 b\bar{X}C_{z(2)}^2, \\
 B_2 &= \lambda(2bQM\bar{X}C_z^2 - 2Q\bar{Y}C_{yz} + 2QPb\bar{X}C_{xz} - 2b\bar{X}C_{xz} - 2b\bar{X}MC_z^2) + Q\theta(Mb\bar{X}C_{z(2)}^2 - 2\bar{Y}C_{yz(2)} + 2b\bar{X}C_{xz(2)}), \\
 B_3 &= \lambda(b\bar{X}C_x^2 - 2\bar{Y}C_{yx} - 2M\bar{Y}C_{yz} + 2Mb\bar{X}C_{xz} + bM^2\bar{X}C_x^2) + \theta(b\bar{X}C_{x(2)}^2 - 2\bar{Y}C_{yx(2)} - 2\bar{Y}MC_{yz(2)} + 2Mb\bar{X}C_{xz(2)} + bM^2\bar{X}C_{x(2)}^2), \\
 A_4 &= A_3 + \theta b\bar{X}C_{x(2)}^2, & B_4 &= B_3 + \theta b\bar{X}C_{x(2)}^2, \\
 A_5 &= A_3 + \lambda b\bar{X}C_{x(2)}^2, & B_5 &= B_3 + \lambda b\bar{X}C_{x(2)}^2, \\
 A_6 &= A_3 + b\bar{X}C_{x(2)}^2(\lambda + \theta) + \eta\rho_{yz}^2 C_y^2, & B_6 &= B_3 + b\bar{X}C_{x(2)}^2(\lambda + \theta) + \eta\rho_{yz}^2 C_y^2, \\
 A_7 &= A_3 + \lambda b\bar{X}C_{x(2)}^2 + \eta\rho_{yz}^2 C_y^2, & B_7 &= B_3 + \lambda b\bar{X}C_{x(2)}^2 + \eta\rho_{yz}^2 C_y^2 \\
 M &= \frac{\lambda(\bar{Y}C_{yz} - b\bar{X}C_{xz}) + \theta(\bar{Y}C_{yz(2)} - b\bar{X}C_{xz(2)})}{b\bar{X}(\lambda C_z^2 + \theta C_{z(2)}^2)}, & Q &= \frac{\lambda C_z^2}{(\lambda C_z^2 + \theta C_{z(2)}^2)}, \\
 P &= \frac{b\bar{X}C_{yz} - \bar{Y}C_{yz} + Mb\bar{X}C_z^2}{b\bar{X}C_z^2 - Qb\bar{X}C_z^2}
 \end{aligned}$$

## 5 An empirical study

To study the performance of the proposed estimator, we use an empirical study given by Khare and Sinha [22]. The description of the population is given below: The data on physical growth of upper socio-economic group of 95 school children of Varanasi under an ICMR study, Department of Pediatrics, B.H.U., during 1983-84 has been taken under study. The first 25% (i. e. 24 children) units have been considered as non-responding units. Here we have taken the study character (y), auxiliary character (x) and the additional auxiliary character (z) are taken as follows:

y: weight (in kg.) of the children.

x: skull circumference (in cm) of the children.

z: chest circumference (in cm) of the children.

The values of the parameters of the characters for the given data are given as follows:

$$\begin{aligned} \bar{Y} &= 19.4968, & \bar{X} &= 51.1726, & \bar{Z} &= 55.8611, & C_y &= 0.15613, \\ C_x &= 0.03006, & C_z &= 0.05711, & C_{y(2)} &= 0.12075, & C_{x(2)} &= 0.032479, \\ C_{z(2)} &= 0.05403, & \rho_{yx} &= 0.328, & \rho_{yz} &= 0.868, & \rho_{xz} &= 0.305, \\ \rho_{yx(2)} &= 0.477, & \rho_{yz(2)} &= 0.729, & \rho_{xz(2)} &= 0.507, & W_2 &= 0.25, \\ N &= 95, & n &= 35. \end{aligned}$$

**Table 1, Relative efficiency (in %) of the estimators with respect to  $\bar{y}^*$  for the fixed values of  $n'$ ,  $n$  and different values of  $k(N=95, n'=70$  and  $n=35$ ).**

Estimators	1/k		
	1/4	1/3	1/2
$\bar{y}^*$	100(0.28597)*	100(0.24638)	100(0.20679)
$T_1$	116(0.24569)	115(0.21490)	112(0.18403)
$T_2$	105(0.27172)	106(0.23213)	107(0.19255)
$T_3$	130(0.21847)	131(0.18789)	132(0.15730)
$T_4$	117(0.24546)	120(0.20588)	124(0.16629)
$T_5$	185(0.15489)	186(0.13268)	188(0.10980)

\*Figures in parenthesis given the MSE(.).

## 6 Conclusion

From Table 1, we observed that the proposed estimator  $T_5$  is more efficient than  $\bar{y}^*$ ,  $T_1$ ,  $T_2$ ,  $T_3$  and  $T_4$  for the different values of  $k$ . The MSE of all estimators decrease the value of  $k$  decrease. The relative efficiency of  $T_2$ ,  $T_3$ ,  $T_4$  and  $T_5$  with respect to  $\bar{y}^*$  increase as  $k$  decrease but relative efficiency of  $T_1$  with respect to  $\bar{y}^*$  decrease as  $k$  decrease. This is due to the fact that with the decrease in the value of  $k$ , the variance of  $\bar{y}^*$  decrease with the faster rate in comparison to  $T_1$ . Finally we observe that the proposed estimator is more effective than the relevant estimators for the different values of  $k$ . Hence we prefer the use of  $T_5$  in practice.

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