

An Exponential Estimator over Regression Estimator Using Two Auxiliary variables

Sachin Malik

Department of Mathematics, SRM University Delhi-NCR, Sonapat, Haryana, India

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Abstract: In the present study, we propose an estimator for population mean when the information is available for two variables under simple random sampling. Expressions for the MSE's of the proposed estimator are derived up to the first degree of approximation. The theoretical conditions have also been verified by a numerical example. It has been shown that the proposed estimator is more efficient than usual regression estimator for two variables.
Keywords: Simple random sampling, auxiliary variable, mean square error, efficiency.

1 Introduction

The use of auxiliary information can increase the precision of an estimator when study variable y is highly correlated with auxiliary variable x . The regression estimator for estimating the unknown population mean

$$t_1 = \bar{y} + b(\bar{X} - \bar{x}) \tag{1}$$

The MSE expression of regression estimator is

$$MSE(t_1) = f_1 \bar{Y}^2 C_y^2 (1 - \rho^2) \tag{2}$$

Where, $f_1 = \frac{N-n}{Nn}$, n is the sample size, N is the population size, (C_y, C_x) are the coefficients of variation of the variates (y, x) respectively.

When there are two auxiliary variables X and Z , the regression estimator of \bar{Y} is

$$t_2 = \bar{y} + b_1(\bar{X} - \bar{x}) + b_2(\bar{Z} - \bar{z}) \tag{3}$$

Where $b_1 = \frac{s_{yx}}{s_x^2}$ and $b_2 = \frac{s_{yz}}{s_z^2}$. Here s_x^2 and s_z^2 are the sample variances respectively, s_{yx} and s_{yz} are the sample covariance respectively. The MSE expression of this estimator is

$$MSE(t_2) = f_1 \bar{Y}^2 C_y^2 (1 - \rho_{yx}^2 - \rho_{yz}^2 + 2\rho_{yx}\rho_{yz}\rho_{xz}) \tag{4}$$

*Corresponding author e-mail: sachinkurava999@gmail.com

2 The proposed Estimator

Following Malik and Singh [1], Bahl and Tuteja [2], we propose the multivariate ratio estimator over regression estimator using information of two auxiliary variables as follows

$$t_p = \bar{y} \exp\left(\alpha_1 \frac{\bar{X} - \bar{x}}{\bar{X} + \bar{x}}\right) \exp\left(\alpha_2 \frac{\bar{Z} - \bar{z}}{\bar{Z} + \bar{z}}\right) + b_1(\bar{X} - \bar{x}) + b_2(\bar{Z} - \bar{z}) \quad (5)$$

To obtain the MSE of t_p to the first degree of approximation, we define

$$e_0 = \frac{\bar{y} - \bar{Y}}{\bar{Y}}, \quad e_1 = \frac{\bar{x} - \bar{X}}{\bar{X}}, \quad e_2 = \frac{\bar{z} - \bar{Z}}{\bar{Z}}$$

Such that, $E(e_i) = 0$; $i = 0, 1, 2$.

$$E(e_0^2) = f_1 C_y^2, \quad E(e_1^2) = f_1 C_x^2, \quad E(e_2^2) = f_1 C_z^2,$$

$$E(e_0 e_1) = f_1 K_{yx} C_x^2, \quad E(e_0 e_2) = f_1 K_{yz} C_z^2, \quad E(e_1 e_2) = f_1 K_{xz} C_z^2,$$

$$K_{yx} = \rho_{yx} \frac{C_y}{C_x}, \quad K_{yz} = \rho_{yz} \frac{C_y}{C_z}, \quad K_{xz} = \rho_{xz} \frac{C_x}{C_z}.$$

Expressing equation (5) in terms of e 's, we have

$$\begin{aligned} t_p &= \bar{Y} (1 + e_0) \left(\exp\left(\alpha_1 \frac{-e_1}{2 + e_1}\right) \exp\left(\alpha_2 \frac{-e_2}{2 + e_2}\right) \right) - b_1 \bar{X} e_1 - b_2 \bar{Z} e_2 \\ &= \bar{Y} \left[1 + e_0 - \frac{\alpha_1 e_1}{2} + \frac{\alpha_1^2 e_1^2}{8} - \frac{\alpha_2 e_2}{2} - \frac{\alpha_1 \alpha_2 e_1 e_2}{4} + \frac{\alpha_2^2 e_2^2}{8} - \frac{\alpha_2 e_0 e_2}{2} - \frac{\alpha_1 e_0 e_1}{2} \right] - b_1 \bar{X} e_1 - b_2 \bar{Z} e_2 \end{aligned} \quad (6)$$

Retaining the term's up to single power of e 's in (6) we have

$$t_p - \bar{Y} = \left\{ \bar{Y} \left[e_0 - \frac{\alpha_1 e_1}{2} - \frac{\alpha_2 e_2}{2} \right] - b_1 \bar{X} e_1 - b_2 \bar{Z} e_2 \right\} \quad (7)$$

Squaring both sides of (7) and then taking expectations, we get the MSE of the estimator t_p up to the first order of approximation, as

$$\begin{aligned} \text{MSE}(t_p) &= f_1 \left\{ \bar{Y}^2 \left[C_y^2 + \frac{\alpha_1^2 C_x^2}{4} + \frac{\alpha_2^2 C_z^2}{4} + \frac{\alpha_1 \alpha_2 k_{xz} C_z^2}{2} - \alpha_1 k_{yx} C_x^2 - \alpha_2 k_{yz} C_z^2 \right] \right. \\ &\quad \left. + B_1^2 \bar{X}^2 C_x^2 + B_2^2 \bar{Z}^2 C_z^2 + 2B_1 B_2 \bar{X} \bar{Z} k_{xz} C_z^2 - 2\bar{Y} \left[B_1 \bar{X} k_{yx} C_x^2 + B_2 \bar{Z} k_{yz} C_z^2 - \frac{\alpha_1 B_1 \bar{X} C_x^2}{2} \right] \right\} \end{aligned}$$

$$\left. \begin{aligned} &-\frac{\alpha_2 B_2 \bar{Z} C_z^2}{2} - \frac{\alpha_1 B_2 \bar{K}_{xz} C_z^2}{2} - \frac{\alpha_2 B_1 \bar{K}_{xz} C_z^2}{2} \end{aligned} \right\} \tag{8}$$

$$MSE(t_p) = A_1 + \alpha_1^2 A_2 + \alpha_2^2 A_3 + \alpha_1 \alpha_2 A_4 - \alpha_1 A_5 - \alpha_2 A_6 \tag{9}$$

$$A_1 = f_1 [\bar{Y}^2 C_y^2 + B_1^2 \bar{X}^2 C_x^2 + B_2^2 \bar{Z}^2 C_z^2 + 2B_1 B_2 \bar{X} \bar{Z} K_{xz} C_z^2 - 2\bar{Y} \bar{X} B_1 K_{yx} C_x^2 - 2\bar{Y} \bar{Z} B_2 K_{xz} C_z^2]$$

$$A_2 = \frac{f_1 \bar{Y}^2 C_x^2}{4}, A_3 = \frac{f_1 \bar{Y}^2 C_z^2}{4}, A_4 = \frac{f_1 \bar{Y}^2 K_{xz} C_z^2}{2}$$

$$A_5 = f_1 [K_{yx} C_x^2 - B_1 \bar{Y} \bar{X} C_x^2 - \bar{Y} \bar{Z} B_2 K_{xz} C_z^2]$$

$$A_6 = f_1 [K_{xz} C_z^2 - B_1 \bar{Y} \bar{X} K_{xz} C_z^2 - \bar{Y} \bar{Z} B_2 C_z^2]$$

Where, $B_1 = \frac{S_{yx}}{S_x^2}$ and $B_2 = \frac{S_{yz}}{S_z^2}$.

Minimising equation (9) with respect to α_1 and α_2 we get the optimum values as

$$\alpha_1 = \frac{A_4 A_6 - 2A_3 A_5}{A_4^2 - 4A_2 A_3} \text{ and } \alpha_2 = \frac{A_4 A_5 - 2A_2 A_6}{A_4^2 - 4A_2 A_3}$$

3 Numerical Illustrations

We have applied the traditional and proposed estimator on the data of apple production amount in 1999 (as interest of variate) and number of apple trees in 1999 (as first auxiliary variate), apple production amount in 1998 (as second auxiliary variate) of 204 villages in Black Sea Region of Turkey (Source: Institute of Statistics, Republic of Turkey). For this data, we have

$N=204, n=50, \bar{Y}=966, \bar{X}=26441, \bar{Z}=1014, S_y=2389.76, S_x=45402.78, S_z=2521.40,$ $S_{xz}=94636084, S_{yx}=77372777, S_{yz}=5684276, \rho_{xz}=0.83, \rho_{yx}=0.71, \rho_{yz}=0.94$
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Table 3.1: MSE values of estimators.

Estimators	\bar{y}	t ₂ (Regression)	t _p (Proposed)
MSE Values	86224.19	62097.28	2853.37

4 Conclusions

In this paper, we have proposed an exponential estimator over regression estimator for estimating unknown population mean of study variable using information on two auxiliary variables. From Table 3.1, we observe that the proposed estimator t_p is best followed by the regression estimator for two auxiliary variables.

Conflict of Interest

The authors declare that they have no conflict of interest.

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